Finding Paths in Sparse Random Graphs Requires Many Queries*

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ABSTRACT: We discuss a new algorithmic type of problem in random graphs studying the minimum number of queries one has to ask about adjacency between pairs of vertices of a random graph $G \sim \mathcal{G}(n,p)$ in order to find a subgraph which possesses some target property with high probability. In this paper we focus on finding long paths in $G \sim \mathcal{G}(n,p)$ when $p = \frac{1+\varepsilon}{n}$ for some fixed constant $\varepsilon > 0$. This random graph is known to have typically linearly long paths.

To have ℓ edges with high probability in $G \sim \mathcal{G}(n, p)$ one clearly needs to query at least $\Omega\left(\frac{\ell}{p}\right)$ pairs of vertices. Can we find a path of length ℓ economically, i.e., by querying roughly that many pairs? We argue that this is not possible and one needs to query significantly more pairs. We prove that any randomised algorithm which finds a path of length $\ell = \Omega\left(\frac{\log\left(\frac{1}{\varepsilon}\right)}{\varepsilon}\right)$ with at least constant probability in $G \sim \mathcal{G}(n, p)$ with $p = \frac{1+\varepsilon}{n}$ must query at least $\Omega\left(\frac{\ell}{p\varepsilon\log\left(\frac{1}{\varepsilon}\right)}\right)$ pairs of vertices. This is tight up to the log $\left(\frac{1}{\varepsilon}\right)$ factor. © 2016 Wiley Periodicals, Inc. Random Struct. Alg., 50, 71–85, 2017

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1. INTRODUCTION

Let \mathcal{P} be a monotone increasing graph property (that is, a property of graphs that cannot be violated by adding edges). Suppose that the edge probability p = p(n) is chosen so that a random graph *G* drawn from the probability space $\mathcal{G}(n, p)$ has \mathcal{P} with high probability (whp). How many queries of the type "is $(i,j) \in E(G)$?" are needed for an adaptive algorithm interacting with the probability space $\mathcal{G}(n, p)$ in order to reveal whp a subgraph $G' \subseteq G$ possessing \mathcal{P} ?

This fairly natural algorithmic setting (see the excellent survey of Frieze and McDiarmid [10] for an extensive coverage of a variety of problems and results in Algorithmic Theory of Random Graphs) has been considered implicitly in several papers on random graphs (e.g. [5, 14]), but apparently has been stated explicitly only in the companion paper [9] of the authors. Notice that in this framework the issue of concern is not the amount of computation required for the algorithm to find a target structure, but rather the amount of its interaction with the underlying probability space.

In the discussion below, we assume some basic familiarity with results about the probability space $\mathcal{G}(n, p)$; the reader is advised to consult monographs [11] and [6] for background on the subject.

In general, given a monotone property \mathcal{P} , what can we expect? If all *n*-vertex graphs belonging to \mathcal{P} have at least *m* edges, then the algorithm should get at least *m* positive answers to hit the target property with the required absolute certainty. This means that the obvious lower bound in this case is at least (1 + o(1))m/p queries. Perhaps one of the simplest graph properties to consider in this respect is connectedness: for any connected graph *G* on *n* vertices a spanning tree can be found after n - 1 queries with positive answers – the algorithm starts with an arbitrary vertex $v \in V(G)$, and each time queries the pairs leaving the current tree until the first edge is found, the tree is then updated by appending this edge. Thus for the regime where $\mathcal{G}(n, p)$ is whp connected (which is when $p(n) \geq \frac{\ln n + \omega(n)}{n}$ with $\lim_{n\to\infty} \omega(n) = 1$), we get an algorithm whp discovering a spanning tree after querying (1 + o(1))n/p pairs of vertices.

A much more challenging problem is that of Hamiltonicity, i.e., of finding a Hamilton cycle. In this case the trivial lower bound translates to *n* positive answers. In [9] we show that this lower bound is tight by providing an adaptive algorithm interacting with the probability space $\mathcal{G}(n, p)$, which whp finds a Hamilton cycle in $G \sim \mathcal{G}(n, p)$ after obtaining only (1 + o(1))n positive answers (provided *p* is above the sharp threshold for Hamiltonicity in $\mathcal{G}(n, p)$).

Yet another positive example is that of uncovering a giant component in the supercritical regime $p = \frac{1+\varepsilon}{n}$. Though this was not the main concern in [14], the second and the third author presented there a very natural adaptive algorithm (essentially performing the Depth First Search (DFS) on a random input $G \sim \mathcal{G}(n, p)$), typically discovering a connected component of size at least $\epsilon n/2$ after querying $\epsilon n^2/2$ vertex pairs.

Upon reviewing these results, the reader may arrive at a conclusion that the above stated trivial lower bound for this type of problems is nearly tight for almost every natural graph property. However, this happens **not** to be the case, and the main qualitative goal of the present paper is to provide such a negative example, including its analysis. Here we focus on the property of containing a path of length ℓ in the supercritical regime in $G \sim \mathcal{G}(n, p)$, that is, when $p = \frac{1+\varepsilon}{n}$ for some fixed constant $\varepsilon > 0$. For this regime, $G \sim \mathcal{G}(n, p)$ is known to contain whp a path of length linear in *n*, due to the classical result of Ajtai, Komlós and Szemerédi [3] (see [14] for a recent simple proof of this fact.) Note that in order to have

 ℓ edges with high probability in $G \sim \mathcal{G}(n, p)$ one needs to query at least $\Omega\left(\frac{\ell}{p}\right)$ pairs of vertices. Can we find a path of length ℓ by asking roughly that many queries, as in the case of Hamiltonicity mentioned above? We show that in this case one actually needs to query significantly more pairs of vertices:

Theorem 1. There exists an absolute constant C > 0 such that the following holds. For every constant $q \in (0, 1)$ there exist $n_0, \varepsilon_0 > 0$ such that for every fixed $\varepsilon \in (0, \varepsilon_0)$ and any $n \ge n_0$ there is no adaptive algorithm which reveals a path of length $\ell \ge \frac{3C}{\varepsilon} \ln\left(\frac{1}{\varepsilon}\right)$ with probability at least q in $G \sim \mathcal{G}(n, p)$, where $p = \frac{1+\varepsilon}{n}$, by querying at most $\frac{q\ell}{8640Cp\varepsilon \ln\left(\frac{1}{\varepsilon}\right)}$ pairs of vertices.

Notice that [14] presents a simple adaptive DFS algorithm, finding a path of length $\frac{1}{5}\varepsilon^2 n$ with probability at least $1 - \exp(\Omega(\varepsilon n))$ in $G \sim \mathcal{G}(n, p)$ after querying only $O(\varepsilon n^2)$ pairs of vertices. In fact, if one uses the same algorithm to find a path of length $\ell \leq \frac{1}{5}\varepsilon^2 n$ in $G \sim \mathcal{G}(n, p)$ then the same argument shows that such a path is found with probability at least $1 - \exp(\Omega(\frac{\ell}{\varepsilon}))$ after querying at most $O(\frac{\ell}{p\varepsilon})$ pairs of vertices. This shows that up to the $\Theta(\log(\frac{1}{\varepsilon}))$ factor, Theorem 1 is tight.

The key ingredient of the proof of Theorem 1 is the following result of independent interest.

Theorem 2. There exist constants $C, \varepsilon_0 > 0$ such that for every fixed $\varepsilon \in (0, \varepsilon_0)$ and $p = \frac{1+\varepsilon}{n}$ we have whp that a graph $G \sim \mathcal{G}(n,p)$ does not contain a set of vertex disjoint paths of lengths at least $\frac{C}{\varepsilon} \ln \left(\frac{1}{\varepsilon}\right)$ whose union covers at least $13\varepsilon^2 n$ vertices.

The rest of this paper is organised as follows. In Section 2 we provide auxiliary lemmas needed for the proofs of Theorem 1 and 2. In Section 3 we prove Theorem 1 assuming Theorem 2. In Section 4 we prove Theorem 2. Finally, in Section 5 we discuss some concluding remarks.

Notation. Our notation is fairly standard. Given a natural number *n* we use [n] to denote the set $\{1, 2, ..., n\}$. Moreover, given a set *V* we use S_V to denote the permutation group of *V* and $\binom{V}{2}$ to denote the set of all (unordered) pairs of elements in *V*.

Given a subset S of the vertex set of a graph G, G[S] denotes the subgraph of G induced by the vertices in S, i.e. the graph with vertex set S whose edges are the ones of G between vertices in S.

A subgraph *P* of the graph *G* is called a *path* if $V(P) = \{v_1, \ldots, v_\ell\}$ and the edges of *P* are $v_1v_2, v_2v_3, \ldots, v_{\ell-1}v_\ell$. We shall oftentimes refer to *P* simply by $v_1v_2 \ldots v_\ell$. We say that such a path *P* has *length* $\ell - 1$ (number of edges) and *size* ℓ (number of vertices).

If *G* is a graph then the 2-*core* of *G* is the maximal induced subgraph of *G* of minimum degree at least 2. If no such subgraph exists then the 2-core of *G* is the empty graph.

Given an ordered set *V* and a real number $p \in [0, 1]$, the binomial random graph model $\mathcal{G}(V, p)$ is a probability space whose ground set consists of all labeled graphs on the vertex set *V*. We can describe the probability distribution of $G \sim \mathcal{G}(V, p)$ by saying that each pair of elements of *V* forms an edge in *G* independently with probability *p*. If V = [n] then we will abuse notation slightly and use $\mathcal{G}(n, p)$ to refer to $\mathcal{G}([n], p)$. Given a property \mathcal{P} (that is, a collection of graphs) and a function $p = p(n) \in [0, 1]$, we say that $G \sim \mathcal{G}(n, p)$ has

 \mathcal{P} with high probability (or whp for brevity) if the probability that $G \in \mathcal{P}$ tends to 1 as *n* tends to infinity.

2. AUXILIARY LEMMAS

2.1. Concentration Inequalities

We need to employ standard bounds on large deviations of random variables. The following well-known lemma due to Chernoff (commonly known as the "Chernoff bound") provides a bound on the upper tail of the Binomial distribution (see e.g. [4, 11]).

Lemma 1. Let $X \sim Bin(n,p)$ and let $\mu = \mathbb{E}[X]$. Then $Pr[X \ge (1+a)\mu] < e^{-\frac{a^2\mu}{3}}$ for any $0 < a < \frac{3}{2}$.

The next lemma is a concentration inequality for the edge exposure martingale in $\mathcal{G}(n, p)$ which follows easily from Theorem 7.4.3 of [4].

Lemma 2. Suppose X is a random variable in the probability space $\mathcal{G}(n,p)$ such that $|X(G) - X(H)| \leq C$ if G and H differ in one edge. Then

$$\Pr\left[|X - \mathbb{E}[X]| > C\alpha \sqrt{n^2 p}\right] \le 2e^{-\frac{\alpha^2}{4}}$$

for any positive $\alpha < 2\sqrt{n^2 p}$.

2.2. Galton-Watson Trees and Paths

A Galton-Watson tree is a random rooted tree, constructed recursively from the root where each node has a random number of children and these random numbers are independent copies of some random variable ξ taking values in $\{0, 1, 2, \ldots\}$. We let \mathcal{T} denote a (random) Galton-Watson tree. We view the children of each node as arriving in some random order, so that \mathcal{T} is an ordered, or plane tree.

We consider the *conditioned Galton-Watson tree* T_t , which is the random tree T conditioned on having exactly *t* vertices. In symbols, $T_t := (T | |T| = t)$, where, for any tree *T*, |T| denotes its number of vertices.

For a rooted tree *T*, the *depth* h(v) of a vertex *v* is its distance to the root (in particular the root has depth 0). We define as usual the *height* of the rooted tree *T* by $H(T) := \max\{h(v) : v \in T\}$. The following lemma which appears in [1] provides essentially optimal uniform sub-Gaussian upper tail bounds on $\frac{H(T_t)}{\sqrt{t}}$ for every offspring distribution ξ with finite variance.

Lemma 3. Suppose that $\mathbb{E}[\xi] = 1$ and $0 < Var[\xi] < \infty$. Then there exist constants C, c > 0 (which may depend on ξ) such that

$$\Pr\left[H(\mathcal{T}_t) \ge h\right] \le C \exp\left(-\frac{ch^2}{t}\right)$$

for all $h \ge 0$ and $t \ge 1$.

As is well known, the distribution of the tree \mathcal{T}_t is not changed if ξ is replaced by another random variable ξ' whose distribution is created from that of ξ by *tilting* or *conjugation* (see e.g. [13]): if for every $k \ge 0$ we have $\Pr[\xi' = k] = c'\mu^k \Pr[\xi = k]$ for some $\mu > 0$ and normalizing constant c'. Thus, we see that Lemma 3 remains true for $\xi \sim \text{Poisson}(\mu)$, with $\mu > 0$, in which case the parameters C, c > 0 are universal constants which do not depend on the parameter μ . It is also well known (see e.g. Section 6.4 of [7]) that if $\xi \sim \text{Poisson}(\mu)$ then \mathcal{T}_t is distributed as a random rooted labelled tree, that is, a tree picked uniformly from the t^{t-1} trees on vertices $\{1, 2, \ldots, t\}$ in which one vertex is declared to be the root. From this we obtain an estimate to be used by us later.

Lemma 4. Given $0 \le \ell \le t$ let $p_{t,\ell}$ denote the proportion of (rooted) labeled trees on t vertices which contain a path of length at least ℓ . There exist constants $C, \varepsilon_0 > 0$ such that for any $\varepsilon \in (0, \varepsilon_0)$ if $\ell = \frac{C}{\varepsilon} \ln \left(\frac{1}{\varepsilon}\right)$ and $t_0 = \frac{15}{\varepsilon^2} \ln \left(\frac{1}{\varepsilon}\right)$ then

$$\sum_{\ell \le t \le t_0} p_{t,\ell} \le \varepsilon^3$$

Proof of Lemma 4. It follows from Lemma 3 and the considerations above that there exist constants C', c' > 0 such that for every $t \le t_0$:

$$p_{t,\ell} \leq C' \exp\left(-\frac{c'\ell^2}{t}\right) \leq C' \exp\left(-\frac{c'\left(\frac{C}{\varepsilon}\ln\left(\frac{1}{\varepsilon}\right)\right)^2}{\frac{15}{\varepsilon^2}\ln\left(\frac{1}{\varepsilon}\right)}\right) = C'\varepsilon^{\frac{c'C^2}{15}}.$$

Thus, if $C > \sqrt{\frac{90}{c'}}$ and if ε_0 is sufficiently small then we see that for any $\varepsilon \in (0, \varepsilon_0)$ and for $t \le t_0$ we have $p_{t,\ell} \le \varepsilon^6$. Using this we conclude that

$$\sum_{\ell \le t \le t_0} p_{t,\ell} \le \varepsilon^6 \cdot t_0 = 15\varepsilon^4 \ln\left(\frac{1}{\varepsilon}\right) \le \varepsilon^3,$$

provided ε_0 is sufficiently small, as claimed.

The next lemma concerns the sizes of Poisson Galton-Watson trees which contain long paths.

Lemma 5. For $\varepsilon > 0$ let $0 < \mu < 1$ be such that $\mu e^{-\mu} = (1 + \varepsilon)e^{-(1+\varepsilon)}$. Given $\ell \ge 1$ consider a Poisson(μ)-Galton-Watson tree T and the random variable

$$T_{\ell} := \begin{cases} |\mathcal{T}| & \text{if } \mathcal{T} \text{ contains a path of length at least } \frac{\ell}{3} \\ 0 & \text{otherwise }, \end{cases}$$

where $|\mathcal{T}|$ denotes the number of vertices of \mathcal{T} . Then there exist constants $C, \varepsilon_0 > 0$ such that for every $\varepsilon \in (0, \varepsilon_0)$ and for $\ell = \frac{C}{\varepsilon} \ln \left(\frac{1}{\varepsilon}\right)$ we have $\mathbb{E}[T_\ell] \le 14\varepsilon^3$ and $Var[T_\ell] \le \frac{8}{\varepsilon^3}$.

Proof. We have

$$\mathbb{E}[T_{\ell}] = \mathbb{E}[\mathbb{E}[T_{\ell} \mid |\mathcal{T}|]] = \sum_{t \ge 1} \Pr[|\mathcal{T}| = t] \cdot \mathbb{E}[T_{\ell} \mid |\mathcal{T}| = t].$$
(1)

It is well-known (see, e.g., Section 6.6 of [7]) that the size of the $Poisson(\mu)$ -Galton-Watson tree \mathcal{T} follows a Borel(μ) distribution, namely,

$$\Pr[|\mathcal{T}| = t] = \frac{t^{t-1} \left(\mu e^{-\mu}\right)^{t}}{\mu \cdot t!}.$$

Moreover, as discussed in the remarks that follow Lemma 3, if we condition a Poisson(μ)-Galton-Watson tree on it having exactly *t* vertices then it is identically distributed to a random rooted labelled tree on *t* vertices. Thus, it follows that $\mathbb{E}[T_{\ell} \mid |\mathcal{T}| = t]$ is equal to $t \cdot p_{t,\frac{\ell}{3}}$, where $p_{t,\frac{\ell}{3}}$ denotes the proportion of rooted labeled trees on *t* vertices which contain a path of length at least $\frac{\ell}{3}$. Hence, setting $t_0 := \frac{15}{\varepsilon^2} \ln(\frac{1}{\varepsilon})$ with foresight, it follows from (1) that

$$\mathbb{E}\left[T_{\ell}\right] = \sum_{t \ge 1} \frac{t^{t-1} \left(\mu e^{-\mu}\right)^{t}}{\mu \cdot t!} \cdot t \cdot p_{t,\frac{\ell}{3}}$$

$$\leq \frac{1}{\mu} \sum_{t \ge \frac{\ell}{3}} \frac{t^{t}}{t!} \cdot (1+\varepsilon)^{t} \cdot e^{-(1+\varepsilon)t} \cdot p_{t,\frac{\ell}{3}}$$

$$\leq 2 \sum_{t \ge \frac{\ell}{3}} e^{-\frac{\varepsilon^{2}}{3}t} \cdot p_{t,\frac{\ell}{3}}$$

$$\leq 2 \cdot \left(\sum_{\frac{\ell}{3} \le t \le t_{0}} p_{t,\frac{\ell}{3}} + \sum_{t \ge t_{0}} e^{-\frac{\varepsilon^{2}}{3}t}\right), \qquad (2)$$

where in the second inequality we used the facts that $\frac{t^{l}}{t!} \leq e^{t}$, that $1 + \varepsilon \leq e^{\varepsilon - \frac{\varepsilon^{2}}{3}}$ (which holds since the first terms of the Taylor series expansion of $\ln(1 + \varepsilon)$ are $\varepsilon - \frac{\varepsilon^{2}}{2}$) and that $\frac{1}{\mu} \leq 2$ provided ε_{0} is chosen sufficiently small. By Lemma 4 there exist constants $C, \varepsilon_{0} > 0$ such that the first sum in (2) is at most ε^{3} . Moreover, the second sum in (2) is

$$\sum_{t \ge t_0} e^{-\frac{\varepsilon^2}{3}t} = e^{-\frac{\varepsilon^2}{3}t_0} \cdot \frac{1}{1 - e^{-\frac{\varepsilon^2}{3}}} \le \varepsilon^5 \cdot \frac{6}{\varepsilon^2} = 6\varepsilon^3,$$
(3)

where we used the fact that $\frac{1}{1-e^{-x}} \leq \frac{2}{x}$ for x > 0 sufficiently small (which holds since the first terms of the Taylor series expansion of e^{-x} are 1-x). Thus, all in all, we conclude that there exist constants $C, \varepsilon_0 > 0$ such that

$$\mathbb{E}\left[T_{\ell}\right] \le 2 \cdot (\varepsilon^3 + 6\varepsilon^3) = 14\varepsilon^3$$

as claimed. Since $|\mathcal{T}| \sim \text{Borel}(\mu)$ it follows that

$$\operatorname{Var}\left[T_{\ell}\right] \leq \mathbb{E}\left[T_{\ell}^{2}\right] \leq \mathbb{E}\left[|\mathcal{T}|^{2}\right] = \frac{1}{(1-\mu)^{3}}.$$

If $\mu \leq 1 - \frac{\varepsilon}{2}$ then we can conclude that $\operatorname{Var}[T_{\ell}] \leq \frac{8}{\varepsilon^3}$, finishing the proof.

It suffices then to show that $\mu \leq 1 - \frac{\varepsilon}{2}$ provided ε_0 is chosen small enough. This is an immediate consequence of the fairly standard estimate in the theory of random graphs

that $\mu = 1 - \varepsilon + O(\varepsilon^2)$ as $\varepsilon \to 0$ (see, e.g. p. 140 of [6]). For the sake of completeness we provide a brief sketch here. Recall that $\mu \in (0, 1)$ is defined as being the solution to $\mu e^{-\mu} = (1 + \varepsilon)e^{-(1+\varepsilon)}$. Let $f : \mathbb{R} \to \mathbb{R}$ denote the function $f(x) = xe^{-x}$. Note that $f'(x) = (1 - x)e^{-x}$, which is strictly positive for $x \in (0, 1)$. This implies that f is strictly increasing in (0, 1) and so, in order to show that $\mu \le 1 - \frac{\varepsilon}{2}$, it is enough to show that $f(1 - \frac{\varepsilon}{2}) \ge (1 + \varepsilon)e^{-(1+\varepsilon)} = f(\mu)$, provided $\varepsilon > 0$ is small enough. Note that:

$$f\left(1-\frac{\varepsilon}{2}\right) = \left(1-\frac{\varepsilon}{2}\right)e^{-\left(1-\frac{\varepsilon}{2}\right)} \ge (1+\varepsilon)e^{-(1+\varepsilon)} \Leftrightarrow \left(1-\frac{\varepsilon}{2}\right)e^{\frac{3\varepsilon}{2}} \ge 1+\varepsilon$$

Since $e^x = \sum_{n=0}^{\infty} \frac{x^n}{n!} \ge 1 + x + \frac{x^2}{2}$ for $x \ge 0$, it follows that:

$$\left(1-\frac{\varepsilon}{2}\right)e^{\frac{3\varepsilon}{2}} \ge \left(1-\frac{\varepsilon}{2}\right)\left(1+\frac{3\varepsilon}{2}+\frac{\left(\frac{3\varepsilon}{2}\right)^2}{2}\right) = 1+\varepsilon+\frac{3\varepsilon^2}{8}-\frac{9\varepsilon^3}{16}$$

The latter is at least $1 + \varepsilon$, if $\varepsilon > 0$ is small enough. Thus, we conclude that $\mu \le 1 - \frac{\varepsilon}{2}$, as claimed.

Lemma 6. Let P = (V, E) be a path of length ℓ and $B \subseteq E$ a set of size $|B| \leq \alpha \ell$, where $\alpha \geq \frac{1}{\ell}$. Let Q denote the graph obtained from P by deleting all the edges in B. Then there exist vertex disjoint subpaths $\{Q^i\}_{i \in J}$ of Q such that each Q^i has length at least $\frac{1}{3\alpha}$ and the subpaths $\{Q^j\}_{i \in J}$ cover at least $(\frac{1}{3} - \alpha) \ell$ vertices of V.

Proof of Lemma 6. Since *P* is a path, *Q* consists of a union of vertex disjoint paths $\{Q^j\}_{j \in [k]}$ for some $k \le |B| + 1 \le \alpha \ell + 1$. Denoting by ℓ_j the length of the path Q^j for $j \in [k]$, note that

$$\sum_{j \in [k]} \ell_j = \ell - |B| \ge (1 - \alpha)\ell.$$
(4)

Moreover, setting $J := \{j \in [k] : \ell_j \ge \frac{1}{3\alpha}\}$ we see that

$$\sum_{j \notin J} \ell_j \le k \cdot \frac{1}{3\alpha} \le \frac{1}{3}\ell + \frac{1}{3\alpha} \le \frac{2}{3}\ell.$$
(5)

Putting (4) and (5) together we get that

$$\sum_{j\in J}\ell_j\geq \left(\frac{1}{3}-\alpha\right)\ell.$$

Thus, it follows that the paths $\{Q^j\}_{j\in J}$ satisfy the desired conditions.

2.3. Properties of Random Graphs

The next lemma provides bounds on the sizes of the largest and second largest connected components of $G \sim \mathcal{G}(n,p)$ as well as the size of its 2-core when $p = \frac{1+\varepsilon}{n}$, where $\varepsilon > 0$ is a small constant. This lemma is a simple consequence of Theorem 5.4 of [11] and Theorem 3 of [15].

Lemma 7. Let $p = \frac{1+\varepsilon}{n}$ where $\varepsilon > 0$ is a constant. Then there exists a constant $\varepsilon_0 > 0$ such that for every $\varepsilon \in (0, \varepsilon_0)$ the following holds whp for $G \sim \mathcal{G}(n, p)$:

- (a) the largest connected component of G has between εn and $3\varepsilon n$ vertices.
- (b) the second largest connected component of G has at most $\frac{20}{c^2} \ln n$ vertices.
- (c) the 2-core of the largest connected component of G has at most $2\varepsilon^2 n$ vertices.

In [8], Ding, Lubetzky and Peres established a complete characterization of the structure of the giant component C_1 of $G \sim \mathcal{G}(n, p)$ in the strictly supercritical regime $(p = \frac{1+\varepsilon}{n})$ with $\varepsilon > 0$ constant). This was achieved by offering a tractable contiguous model \tilde{C}_1 , i.e. a model such that every graph property that is satisfied by \tilde{C}_1 whp is also satisfied by C_1 whp. In their model, \tilde{C}_1 consists of a 2-core $\tilde{C}_1^{(2)}$ where one attaches to each vertex of $\tilde{C}_1^{(2)}$ one independent Poisson(μ)-Galton-Watson tree (where $0 < \mu < 1$ is such that $\mu e^{-\mu} = (1 + \varepsilon)e^{-(1+\varepsilon)}$. In light of this, any graph property that is satisfied whp by the disjoint union of $|\tilde{C}_1^{(2)}|$ independent Poisson(μ)-Galton-Watson trees must also be satisfied whp by $C_1 \setminus C_1^{(2)}$, the graph obtained from the giant component C_1 by removing the edges of its 2-core $C_1^{(2)}$. As one would expect, the random variable $|\tilde{C}_1^{(2)}|$ is tightly concentrated around its expectation, which agrees with the expected size of the 2-core $C_1^{(2)}$ of C_1 . By (c) of Lemma 7 this at most $2\varepsilon^2 n$. The next technical lemma which will be useful in the proof of Theorem 2 follows from the considerations above.

Lemma 8. Let C_1 denote the largest connected component of $G \sim \mathcal{G}(n,p)$ for $p = \frac{1+\varepsilon}{n}$, where $\varepsilon > 0$ is fixed, let $C_1^{(2)}$ denote its 2-core and let $C_1 \setminus C_1^{(2)}$ denote the graph obtained from C_1 by removing the edges in $C_1^{(2)}$. Let $0 < \mu < 1$ be such that $\mu e^{-\mu} = (1 + \varepsilon)e^{-(1+\varepsilon)}$ and consider $2\varepsilon^2 n$ independent Poisson(μ)-Galton-Watson trees $T_1, \ldots, T_{2\varepsilon^2 n}$. Then, for every ℓ and m (which might depend on n) if whp the disjoint union of $T_1, \ldots, T_{2\varepsilon^2 n}$ does not contain a set of vertex disjoint paths of length at least ℓ covering at least m vertices then the same holds whp for $C_1 \setminus C_1^{(2)}$.

3. PROOF OF THEOREM 1

We start this section by repeating the statement of Theorem 1 for the reader's convenience.

Theorem 1. There exists an absolute constant C > 0 such that the following holds. For every constant $q \in (0, 1)$ there exist $n_0, \varepsilon_0 > 0$ such that for every fixed $\varepsilon \in (0, \varepsilon_0)$ and any $n \ge n_0$ there is no adaptive algorithm which reveals a path of length $\ell \ge \frac{3C}{\varepsilon} \ln\left(\frac{1}{\varepsilon}\right)$ with probability at least q in $G \sim \mathcal{G}(n, p)$, where $p = \frac{1+\varepsilon}{n}$, by querying at most $\frac{q\ell}{8640Cp\varepsilon \ln\left(\frac{1}{\varepsilon}\right)}$ pairs of vertices.

Proof of Theorem 1. Suppose Alg is an adaptive algorithm which with probability at least q finds a path of length ℓ in $G \sim \mathcal{G}(n, p)$, where $p = \frac{1+\varepsilon}{n}$, after querying at most $\frac{q\ell}{8640Cp\varepsilon\ln\left(\frac{1}{\varepsilon}\right)}$ pairs of vertices. We consider implicitly that Alg takes an *ordered* vertex set as part of its input. We shall assume henceforth that n, C > 0 are sufficiently large and $\varepsilon > 0$ is sufficiently small in order to obtain a contradiction. Note that, by restricting Alg to a set of n vertices, we get an algorithm which for any $n' \ge n$ with probability at least q finds in

 $G' \sim \mathcal{G}(n', p)$ a path of length ℓ after querying at most $\frac{q\ell}{8640C_{p\epsilon}\ln\left(\frac{1}{\epsilon}\right)}$ pairs of vertices. We shall abuse notation slightly and call Alg to all these algorithms.

Define $n' := \left(1 + \frac{720\varepsilon^2}{q}\right)n$, $V_0 := [n']$, $I_0 := \emptyset$ and $s := \frac{720\varepsilon^2 n}{q(\ell+1)}$. For i = 1, ..., s do the following:

• Apply Alg to $G_{i-1} \sim \mathcal{G}(V_{i-1}, p)$, where the vertices in V_{i-1} are permuted according to a permutation $\pi_i \in S_{V_{i-1}}$ chosen uniformly at random. Let L_i be the graph of all pairs of vertices queried and let $K_i \subseteq L_i$ be the graph of edges present. By the algorithm we know that L_i has at most $\frac{q\ell}{8640Cp\epsilon \ln(\frac{1}{\epsilon})}$ edges. If K_i contains a path of length ℓ then let P_i be one such path, define $V_i := V_{i-1} \setminus V(P_i)$ and set $I_i := I_{i-1} \cup \{i\}$. Otherwise, set $V_i := V_{i-1}$ and $I_i := I_{i-1}$.

Observe that $|V_s| \ge n' - (\ell + 1)s = \left(1 + \frac{720\varepsilon^2}{q}\right)n - \frac{720\varepsilon^2}{q}n = n$ and so we can indeed apply Alg to V_{i-1} for any $i \in [s]$. We define a random graph H with vertex set V_0 in the following way. For every pair of vertices $\{u, v\} \subseteq V_0$ if $\{u, v\} \in E(L_i)$ for some $i \in [s]$ then let i_0 be the smallest such index and set $\{u, v\}$ as an edge of H if and only if $\{u, v\} \in E(K_{i_0})$. Consider all the other pairs $\{u, v\} \subseteq V_0$ as non-edges of H. From the procedure above it follows that for every $\{u, v\} \subseteq V_0$ we have independently that

$$\Pr\left[\{u,v\}\in E(H)\right] \le p = \frac{1+\varepsilon}{n} = \frac{1+\varepsilon}{n'} \cdot \frac{n'}{n} = \frac{(1+\varepsilon)\left(1+\frac{720\varepsilon^2}{q}\right)}{n'} \le \frac{1+2\varepsilon}{n'},$$

provided $\varepsilon \leq \frac{q}{1440}$. Thus, the graph *H* can be viewed as a subgraph of a graph sampled from $\mathcal{G}\left(n', \frac{1+2\varepsilon}{n'}\right)$. In particular, if with probability at least $\frac{q^2}{4}$ the graph *H* contains a set of vertex disjoint paths of length at least $\frac{C}{\varepsilon} \ln\left(\frac{1}{\varepsilon}\right)$ which cover at least $52\varepsilon^2 n'$ vertices then the same must also hold with probability at least $\frac{q^2}{4}$ in $\mathcal{G}\left(n', \frac{1+2\varepsilon}{n'}\right)$. However, this would contradict Theorem 2 and so it suffices to prove the following claim:

Claim. With probability at least $\frac{q^2}{4}$ the graph *H* contains a set of vertex disjoint paths of length at least $\frac{C}{\varepsilon} \ln\left(\frac{1}{\varepsilon}\right)$ which cover at least $52\varepsilon^2 n'$ vertices of V_0 .

Define for each $i \in I_s$ the graph H_i with vertex set V_{i-1} and edge set $\left(\bigcup_{j=1}^{i-1} E(L_j)\right) \cap \binom{V_{i-1}}{2}$ and note that

$$|E(H_i)| \le s \cdot \frac{q\ell}{8640Cp\varepsilon \ln\left(\frac{1}{\varepsilon}\right)} \le \frac{\varepsilon n^2}{12C \ln\left(\frac{1}{\varepsilon}\right)(1+\varepsilon)} \le \frac{\varepsilon}{6C \ln\left(\frac{1}{\varepsilon}\right)} \cdot \binom{|V_{i-1}|}{2}.$$
 (6)

Observe that for each $i \in I_s$ the set $V_{i-1} \setminus V_i$ consists of the vertex set of a path P_i in the graph K_i . For each such i set $B_i := E(P_i) \cap E(H_i)$ and let Q_i denote the graph obtained from P_i by deleting all the edges in B_i . Note crucially that $E(Q_i) \subseteq E(H)$ and that the graphs $\{Q_i\}_{i \in I_s}$ are vertex disjoint.

Consider now the set $I := \left\{ i \in I_s : |B_i| \le \frac{\varepsilon}{3C \ln\left(\frac{1}{\varepsilon}\right)} \ell \right\}$. By Lemma 6 it follows that for any $i \in I$ there exist vertex disjoint subpaths $\{Q_i^j\}_{j \in J_i}$ of Q_i each of length at least $\frac{C}{\varepsilon} \ln\left(\frac{1}{\varepsilon}\right)$

which cover at least $\left(\frac{1}{3} - \frac{\varepsilon}{3C\ln\left(\frac{1}{\varepsilon}\right)}\right)\ell \ge \frac{1}{4}(\ell+1)$ vertices of $V(Q_i)$. Thus, if $|I| \ge \frac{1}{3}sq$ then $\{Q_i^j\}_{i\in I, j\in J_i}$ forms a collection of vertex disjoint paths in H of length at least $\frac{C}{\varepsilon}\ln\left(\frac{1}{\varepsilon}\right)$ which cover at least $\frac{1}{4}(\ell+1) \cdot \frac{1}{3}sq = 60\varepsilon^2 n \ge 52\varepsilon^2 n'$ vertices of V_0 . It suffices to show then that with probability at least $\frac{q^2}{4}$ we have $|I| \ge \frac{1}{3}sq$.

Let $I' := [s] \setminus I$ and note that for every $i \in [s]$ we have

$$\Pr\left[i \in I'\right] = \Pr\left[i \notin I_s\right] + \Pr\left[i \in I' \mid i \in I_s\right] \cdot \Pr\left[i \in I_s\right].$$
(7)

It is clear from the procedure above that for each $i \in [s]$ we have $\Pr[i \in I_s] \ge q$. Note also crucially that, provided $i \in I_s$, the path P_i is a randomly mapped path of length ℓ on the vertex set V_{i-1} . Indeed, this happens because before the *i*-th application of Alg we permuted the vertices of V_{i-1} according to a permutation $\pi_i \in S_{V_{i-1}}$ chosen uniformly at random. Thus, by conditioning on the event that $i \in I_s$, on any possible graph H_i satisfying (6) and on the path $\pi_i^{-1}(P_i)$, we have for any $e \in E(\pi_i^{-1}(P_i))$:

$$\Pr\left[\pi_i(e) \in E(H_i)\right] \leq \frac{\varepsilon}{6C \ln\left(\frac{1}{\varepsilon}\right)},\,$$

and so, by linearity of expectation it follows that:

$$\mathbb{E}\left[|E(P_i) \cap E(H_i)|\right] \leq \frac{\varepsilon}{6C \ln\left(\frac{1}{\varepsilon}\right)} \ell.$$

Thus, by Markov's inequality (see, e.g., [4]) we get that

$$\Pr\left[i\in I'\mid i\in I_s\right]\leq \frac{1}{2}\,,$$

and so by Eq. (7) we see that for any $i \in [s]$ we have $\Pr\left[i \in I'\right] \le 1 - \frac{1}{2} \Pr[i \in I_s] \le 1 - \frac{q}{2}$. It follows then by linearity of expectation that $\mathbb{E}\left[|I'|\right] \le s\left(1 - \frac{q}{2}\right)$. Hence, again by Markov's inequality we conclude that

$$\Pr\left[|I'| \ge \frac{s}{1+\frac{q}{2}}\right] \le 1 - \frac{q^2}{4} \text{ , which implies } \frac{q^2}{4} \le \Pr\left[|I| \ge \frac{sq}{2+q}\right] \le \Pr\left[|I| \ge \frac{sq}{3}\right].$$

This completes the proof.

4. PROOF OF THEOREM 2

Theorem 2. There exist constants $C, \varepsilon_0 > 0$ such that for every fixed $\varepsilon \in (0, \varepsilon_0)$ we have whp that $G \sim \mathcal{G}\left(n, \frac{1+\varepsilon}{n}\right)$ does not contain a set of vertex disjoint paths of lengths at least $\frac{C}{\varepsilon} \ln\left(\frac{1}{\varepsilon}\right)$ whose union covers at least $13\varepsilon^2 n$ vertices.

Proof of Theorem 2. Let $G \sim \mathcal{G}(n,p)$ where $p = \frac{1+\varepsilon}{n}$. Let \mathcal{C}_1 denote the largest connected component of G, let $\mathcal{C}_1^{(2)}$ denote the 2-core of \mathcal{C}_1 and let $\mathcal{C}_1 \setminus \mathcal{C}_1^{(2)}$ denote the graph obtained from \mathcal{C}_1 by deleting the edges in $\mathcal{C}_1^{(2)}$. For $\ell \geq 1$ consider the following random variables:

- X_{ℓ} = number of vertices which belong to connected components of *G* of size at most $\frac{20}{c^2} \ln n$ containing a path of length at least ℓ .
- Y_{ℓ} = maximum number of vertices covered by vertex disjoint paths of length at least ℓ in C_1 .
- Z_{ℓ} = maximum number of vertices covered by vertex disjoint paths of length at least $\frac{\ell}{3}$ in $C_1 \setminus C_1^{(2)}$.

By (*b*) of Lemma 7 it follows that whp $X_{\ell} + Y_{\ell}$ is an upper bound on the maximum number of vertices of *G* covered by vertex disjoint paths of length at least ℓ . Note that we may assume that all the paths considered have size at most 2ℓ by splitting larger paths into several paths of length at least ℓ . Moreover, if *P* is a path of length at least ℓ in C_1 then, since $C_1 \setminus C_1^{(2)}$ consists of a disjoint union of trees, there must exist a subpath *P'* of the path *P* with at least $\frac{|P|}{3} \ge \frac{\ell}{3}$ vertices which lies in $C_1^{(2)}$ or in $C_1 \setminus C_1^{(2)}$. Since $|P| \le 6|P'|$ it follows that $Y_{\ell} \le 6|C_1^{(2)}| + 6Z_{\ell}$.

By (c) of Lemma 7 we know that whp $|\mathcal{C}_1^{(2)}| \leq 2\varepsilon^2 n$, provided ε_0 is chosen small enough. It suffices then to show that there exist constants $C, \varepsilon_0 > 0$ such that for every fixed $\varepsilon \in (0, \varepsilon_0)$ and for $\ell := \frac{C}{\varepsilon} \ln \left(\frac{1}{\varepsilon}\right)$ we have whp that

$$X_{\ell} < \varepsilon^3 n$$
 and $Z_{\ell} < 29\varepsilon^5 n$.

since in that case we have whp that the maximum number of vertices of G covered by vertex disjoint paths of length at least ℓ is at most

$$X_{\ell} + Y_{\ell} \le X_{\ell} + 6|\mathcal{C}_{1}^{(2)}| + 6Z_{\ell} < \varepsilon^{3}n + 6 \cdot 2\varepsilon^{2}n + 6 \cdot 29\varepsilon^{5}n \le 13\varepsilon^{2}n.$$

provided ε_0 is chosen sufficiently small. Lemmas 9 and 10 complete the proof.

Lemma 9. There exist constants $C, \varepsilon_0 > 0$ such that for every fixed $\varepsilon \in (0, \varepsilon_0)$ and for $\ell := \frac{C}{\varepsilon} \ln \left(\frac{1}{\varepsilon}\right)$ we have $X_{\ell} < \varepsilon^3 n$ whp.

Proof of Lemma 9. Given a set $S \subseteq [n]$ of size t, let $S_{\ell}(S)$ (resp. $\mathcal{T}_{\ell}(S)$) denote the set of possible connected graphs (resp. spanning trees) on the vertex set S which contain a path of length at least ℓ . Let X_S denote the indicator random variable of the event that $G[S] \in S_{\ell}(S)$ and that there are no edges in G between S and $[n] \setminus S$. Note that $G[S] \in S_{\ell}(S)$ if and only if there exists $T \in \mathcal{T}_{\ell}(S)$ such that $T \subseteq G[S]$. Thus, by the union bound we have

$$\mathbb{E}\left[X_{S}\right] \leq \left|\mathcal{T}_{\ell}(S)\right| \cdot p^{t-1} \cdot (1-p)^{t(n-t)} \tag{8}$$

where the first term accounts for taking a union bound over all $T \in \mathcal{T}_{\ell}(S)$, the second term accounts for the probability that the edges in *T* are present in *G*[*S*] and the last term accounts for the probability that none of the edges between *S* and $[n] \setminus S$ are present in *G*. Note that $|\mathcal{T}_{\ell}(S)|$ does not depend on the set *S* and is equal to the number of labeled trees on *t* vertices which contain a path of length at least ℓ . More specifically, if $p_{t,\ell}$ denotes the proportion of labeled trees on *t* vertices which contain a path of length at path of length at least ℓ . More specifically, if $p_{t,\ell}$ denotes the proportion of labeled trees on *t* vertices which contain a path of length at least ℓ , then $|\mathcal{T}_{\ell}(S)| = p_{t,\ell} \cdot t^{t-2}$. Observe now that the random variable X_{ℓ} satisfies the following:

$$X_{\ell} \leq \sum_{t=\ell}^{\frac{20}{\varepsilon^2} \ln n} \sum_{S \in \binom{[n]}{t}} t \cdot X_S.$$

We claim that for $\ell := \frac{C}{\varepsilon} \ln(\frac{1}{\varepsilon})$, where C > 0 is a large constant, and for some constant $\varepsilon_0 > 0$, if $\varepsilon \in (0, \varepsilon_0)$ is fixed then $\Pr[X_{\ell} \ge \varepsilon^3 n] = o(1)$. To prove this claim we start by estimating $\mathbb{E}[X_{\ell}]$. Setting $t_0 := \frac{15}{\varepsilon^2} \ln(\frac{1}{\varepsilon})$, we have by the linearity of expectation and by (8) that if ε_0 is sufficiently small then:

$$\mathbb{E}[X_{\ell}] \leq \sum_{t=\ell}^{\frac{20}{\varepsilon^2} \ln n} t \cdot {\binom{n}{t}} \cdot p_{t,\ell} \cdot t^{t-2} \cdot p^{t-1} \cdot (1-p)^{t(n-t)}$$

$$\leq \sum_{t=\ell}^{\frac{20}{\varepsilon^2} \ln n} t \cdot \left(\frac{en}{t}\right)^t \cdot p_{t,\ell} \cdot t^{t-2} \cdot \left(\frac{1+\varepsilon}{n}\right)^{t-1} \left(1-\frac{1+\varepsilon}{n}\right)^{t(n-t)}$$

$$\leq \sum_{t=\ell}^{\frac{20}{\varepsilon^2} \ln n} e^t \cdot t^{-1} \cdot n \cdot p_{t,\ell} \cdot \frac{e^{\varepsilon t-\frac{\varepsilon^2}{3}t}}{1+\varepsilon} \cdot e^{-(1+\varepsilon)t+\frac{(1+\varepsilon)t^2}{n}}$$

$$\leq \frac{(1+o(1))n}{\ell(1+\varepsilon)} \cdot \sum_{t\geq\ell} p_{t,\ell} \cdot e^{-\frac{\varepsilon^2}{3}t}$$

$$\leq \frac{n}{14} \cdot \left(\sum_{\ell\leq t\leq t_0} p_{t,\ell} + \sum_{t\geq t_0} e^{-\frac{\varepsilon^2}{3}t}\right)$$
(9)

where in the third inequality we used the fact that $(1 + \varepsilon)^t \le e^{\varepsilon t - \frac{\varepsilon^2}{3}t}$ for sufficiently small $\varepsilon > 0$. By Lemma 4 there exist constants $C, \varepsilon_0 > 0$ such that the first sum in (9) is at most ε^3 . Moreover, by (3) the second sum in (9) is at most $6\varepsilon^3$. Thus, all in all, we conclude that there exist constants $C, \varepsilon_0 > 0$ such that

$$\mathbb{E}[X_{\ell}] \leq \frac{n}{14} \cdot (\varepsilon^3 + 6\varepsilon^3) = \frac{\varepsilon^3 n}{2}$$

Note that if *G* and *H* differ in precisely one edge then $|X_{\ell}(G) - X_{\ell}(H)| \le \frac{40}{\varepsilon^2} \ln n$ because one edge affects at most two connected components of size at most $\frac{20}{\varepsilon^2} \ln n$. Thus, by Lemma 2 it follows that

$$\Pr\left[X_{\ell} > \varepsilon^{3}n\right] \le \Pr\left[|X_{\ell} - \mathbb{E}[X_{\ell}]| > \frac{\varepsilon^{3}n}{2}\right] \le e^{-\Omega\left(\frac{n}{(\ln n)^{2}}\right)} = o(1).$$

Remark. An alernative approach to the proof of Lemma 9 would be to invoke the so called symmetry rule (see, e.g., Chapter 5.6 of [11]), postulating that in the supercritical regime $p = \frac{1+\varepsilon}{n}$, the subgraph of $G \sim \mathcal{G}(n, p)$ outside the giant component behaves typically as a random graph with subcritical edge probability. One can then estimate the likely contribution of paths of length at least $\ell = \frac{C}{\varepsilon} \ln \left(\frac{1}{\varepsilon}\right)$ coming from the small components to the total volume of vertex disjoint paths of length at least ℓ and to show it to be $O(\varepsilon^2 n)$ whp, using a direct first moment argument. Since we still need to treat the paths residing in the giant component outside the 2-core (the random variable Z_{ℓ}), we chose to adopt a unified approach using the machinery of Galton-Watson trees developed in Section 2.2, and to apply it here as well.

Lemma 10. There exist constants $C, \varepsilon_0 > 0$ such that for every fixed $\varepsilon \in (0, \varepsilon_0)$ and for $\ell := \frac{C}{\varepsilon} \ln \left(\frac{1}{\varepsilon}\right)$ we have $Z_{\ell} < 29\varepsilon^5 n$ whp.

Proof of Lemma 10. Recall that Z_{ℓ} counts the maximum number of vertices covered by vertex disjoint paths of length at least $\frac{\ell}{3}$ in $C_1 \setminus C_1^{(2)}$. Let $0 < \mu < 1$ be such that $\mu e^{-\mu} = (1+\varepsilon)e^{-(1+\varepsilon)}$ and consider $2\varepsilon^2 n$ independent Poisson(μ)-Galton-Watson trees $\mathcal{T}_1, \ldots, \mathcal{T}_{2\varepsilon^2 n}$. By Lemma 8 it suffices for our purposes to show that whp the maximum number of vertices covered by vertex disjoint paths of length at least $\frac{\ell}{3}$ in the disjoint union of $\mathcal{T}_1, \ldots, \mathcal{T}_{2\varepsilon^2 n}$ is less than $29\varepsilon^5 n$, for appropriate $C, \varepsilon_0 > 0$.

For each $1 \le i \le 2\varepsilon^2 n$ consider the following random variable:

$$T_{i,\ell} := \begin{cases} |\mathcal{T}_i| & \text{if } \mathcal{T}_i \text{ contains a path of length at least } \frac{\ell}{3} \\ 0 & \text{otherwise} \end{cases}$$

and set $T_{\ell} = \sum_{i=1}^{2\varepsilon^2 n} T_{i,\ell}$. Clearly T_{ℓ} is an upperbound on the maximum number of vertices covered by vertex disjoint paths of length at least $\frac{\ell}{3}$ in in the disjoint union of $\mathcal{T}_1, \ldots, \mathcal{T}_{2\varepsilon^2 n}$. To finish the proof, we show that whp $T_{\ell} < 29\varepsilon^5 n$, provided $C, \varepsilon_0 > 0$ are chosen appropriately.

By Lemma 5 we know that there exist constants $C, \varepsilon_0 > 0$ such that for every $\varepsilon \in (0, \varepsilon_0)$ and for $\ell = \frac{C}{\varepsilon} \ln \left(\frac{1}{\varepsilon}\right)$ we have $\mathbb{E}\left[T_{i,\ell}\right] \le 14\varepsilon^3$ and $\operatorname{Var}\left[T_{i,\ell}\right] \le \frac{8}{\varepsilon^3}$. Thus, since the random variables $T_{i,\ell}$ are independent, we have that

$$\mathbb{E}\left[T_{\ell}\right] \le 14\varepsilon^{3} \cdot 2\varepsilon^{2}n = 28\varepsilon^{5}n \quad \text{and} \quad \operatorname{Var}\left[T_{\ell}\right] \le \frac{8}{\varepsilon^{3}} \cdot 2\varepsilon^{2}n = \frac{16n}{\varepsilon}.$$

Thus, by Chebyshev's Inequality (see, e.g., [4]) we conclude that

$$\Pr\left[T_{\ell} \ge 29\varepsilon^{5}n\right] \le \Pr\left[|T_{\ell} - \mathbb{E}\left[T_{\ell}\right]| \ge \varepsilon^{5}n\right] \le \frac{\operatorname{Var}\left[T_{\ell}\right]}{\varepsilon^{10}n^{2}} \le \frac{16}{\varepsilon^{11}n} = o(1).$$

5. CONCLUDING REMARKS

We have shown that in order to find a path of length $\ell = \Omega\left(\frac{\log\left(\frac{1}{\varepsilon}\right)}{\varepsilon}\right)$ in $G \sim \mathcal{G}(n,p)$ with at least some constant probability, where $p = \frac{1+\varepsilon}{n}$ with $\varepsilon > 0$ fixed, one needs to query at least $\Omega\left(\frac{\ell}{p\varepsilon\log\left(\frac{1}{\varepsilon}\right)}\right)$ pairs of vertices. This is close to best possible since a randomised depth first search algorithm from [14] finds whp a path of length ℓ after querying at most $O\left(\frac{\ell}{p\varepsilon}\right)$ pairs of vertices. A natural question, which remains open, is to close the gap between these bounds. We believe that every adaptive algorithm which reveals whp a path of length ℓ in $G \sim \mathcal{G}(n,p)$, where $p = \frac{1+\varepsilon}{n}$ with $\varepsilon > 0$ fixed, has to query $\Omega\left(\frac{\ell}{p\varepsilon}\right)$ pairs of vertices.

Recall that, to prove our main result, in Theorem 2 we bounded the total number of vertices covered by vertex disjoint paths of size at least $\Omega\left(\frac{1}{\varepsilon}\log\left(\frac{1}{\varepsilon}\right)\right)$ in a typical graph sampled from $\mathcal{G}(n,p), p = \frac{1+\varepsilon}{n}$, by $O\left(\varepsilon^2 n\right)$. Since a graph $G \sim \mathcal{G}(n,p)$ contains whp a path of length $\Theta(\varepsilon^2 n)$ (see e.g. [11]), this is best possible up to a multiplicative constant. If one can show that a similar statement holds for paths of length $\Omega\left(\frac{1}{\varepsilon}\right)$ then one can modify our proof to obtain a $\Omega\left(\frac{\ell}{p\varepsilon}\right)$ bound in Theorem 1.

In the proof of Theorem 2 we needed to bound the number of vertices covered by vertex disjoint paths of a prescribed length ℓ in a random tree of fixed size *t* (Lemma 5). Our estimate was a bit wasteful because for trees which contained a path of length ℓ we used their total number of vertices *t* instead of the number of vertices covered by vertex disjoint

paths of length ℓ , which is most likely significantly smaller. A way to fix this is to obtain good bounds for the following question:

Question. Given $a = a(t) \in \mathbb{N}$ and $b = b(t) \in \mathbb{N}$ what is the probability that a random tree on t vertices contains b vertex disjoint paths, each of length at least a?

Note that, since the diameter of a random tree on t vertices is whp $\Theta(\sqrt{t})$ (see e.g. [1]), the only interesting regime is when $ab \ge C\sqrt{t}$ for some constant C > 0. Moreover, by splitting paths of length larger than 2a into smaller subpaths of length at least a, we may consider only paths of length between a and 2a.

One possible approach to this problem would be through a nice argument of Joyal ([12], see also [2]). It shows that a random tree \mathcal{T} on t vertices can be obtained from a random map $f:[t] \to [t]$ as follows. First we create the directed graph D (possibly with loops) on vertex set [t] with edges $i \to f(i)$ for each $i \in [t]$. Then we look at a maximal set of vertices $M = \{i_1, \ldots, i_m\} \subseteq [t]$ such that $f|_M$ is a permutation. We remove the directed edges inside *M* and replace them by the path $f(i_1) \rightarrow f(i_2) \rightarrow \ldots \rightarrow f(i_m)$ (where $i_1 < i_2 < \ldots < i_m$). By ignoring the orientations of the edges we obtain the desired tree \mathcal{T} . Note that, since the vertices in M form a path in T, we must have $|M| = O(\sqrt{t})$ whp. Moreover, if we have a path P in T then a moment's thought reveals that either P has at least $\frac{|V(P)|}{3}$ vertices in M or there are $\frac{|V(P)|}{3}$ vertices of P which form a directed path in D. Thus, it follows that if we have a collection of *b* vertex disjoint paths in \mathcal{T} each of length between *a* and 2*a* then *D* contains a collection of vertex disjoint directed paths each of length between $\frac{a+1}{3}$ and 2*a* covering at least $\frac{(a+1)b}{3} - |M|$ vertices. Since $|M| = O(\sqrt{t})$ whp and since we are interested only in the case when $ab \ge C\sqrt{t}$ for some large constant C > 0, it follows that in that case we have, say, at least $\frac{b}{10}$ such paths. Thus, up to changing a and b by constant multiplicative factors, it is enough to estimate the probability that the directed graph D obtained from a random map $f:[t] \to [t]$ contains at least b vertex disjoint directed paths, each of length (at least) a.

We can give a simple upper bound on this probability by taking the union bound over all collections of b vertex disjoint directed paths of length a. This shows that the probability that we want to estimate is at most

$$\frac{t!}{(t-(a+1)b)!b!}\left(\frac{1}{t}\right)^{ab} = \frac{t^b}{b!}\prod_{i=1}^{(a+1)b-1}\left(1-\frac{i}{t}\right) \le e^{b+b\ln(t/b)-\binom{(a+1)b}{2}/t}.$$

Unfortunately, this upper bound is not strong enough to allow us to prove Theorem 2 for paths of length at least $\Omega\left(\frac{1}{\varepsilon}\right)$ because when *b* is roughly a constant and *a* is close to \sqrt{t} the positive term $b \ln (t/b)$ in the exponent is much larger than the negative term $\binom{(a+1)b}{2}/t$. Thus, it would be nice to obtain tighter bounds for the probability in question.

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