

Risk Aggregation under Dependence Uncertainty

Challenges in Theory and Practice

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Outline

- 1 Framework
- 2 VaR and ES Bounds
- 3 Asymptotic Equivalence
- 4 Challenges
- 5 References

Fundamental problem in Finance/Insurance

- Risk factors: $\mathbf{X} = (X_1, \dots, X_d)$
- Model assumption: $X_i \sim F_i, F_i$ known, $i = 1, \dots, d$
- A financial position $\Psi(\mathbf{X})$
- A risk measure/pricing function: $\rho(\Psi(\mathbf{X}))$

Calculate $\rho(\Psi(\mathbf{X}))$

Calculating $\rho(\Psi(\mathbf{X}))$

Example:

- $\Psi(\mathbf{X}) = \sum_{i=1}^d X_i$
- $\rho = \text{VaR}_p$ or $\rho = \text{ES}_p$

Challenge:

- We need a *joint* model for the random vector \mathbf{X}
- Joint models are hard to get by

We will focus on the above special choices of Ψ and ρ .

VaR and ES

VaR_p(X)

For $p \in (0, 1)$,

$$\text{VaR}_p(X) = F_X^{-1}(p) = \inf\{x \in \mathbb{R} : F_X(x) \geq p\}$$

ES_p(X)

For $p \in (0, 1)$,

$$\text{ES}_p(X) = \frac{1}{1-p} \int_p^1 \text{VaR}_q(X) dq \stackrel{(F \text{ cont.})}{=} \mathbb{E}[X | X > \text{VaR}_p(X)]$$

VaR and ES

A related quantity **Left-tail-ES**:

$LES_p(X)$

For $p \in (0, 1)$,

$$LES_p(X) = \frac{1}{p} \int_0^p \text{VaR}_q(X) dq = -\text{ES}_{1-p}(-X)$$

Fréchet problem

Denote

$$\mathcal{S}_d = \mathcal{S}_d(F_1, \dots, F_d) = \left\{ \sum_{i=1}^d X_i : X_i \sim F_i, i = 1, \dots, d \right\}$$

- Every element in \mathcal{S}_d is a possible risk position.
- Determination of \mathcal{S}_d : very challenging.
 - Think about $\mathcal{S}_2(U[0, 1], U[0, 1])$... open question!

Worst- and best-values of VaR and ES

The Fréchet (unconstrained) problems for VaR_p

$$\overline{\text{VaR}}_p(S_d) = \sup\{\text{VaR}_p(S) : S \in \mathcal{S}_d(F_1, \dots, F_d)\},$$

$$\underline{\text{VaR}}_p(S_d) = \inf\{\text{VaR}_p(S) : S \in \mathcal{S}_d(F_1, \dots, F_d)\}.$$

Same notation for ES_p and LES_p .

Worst- and best-values of VaR and ES

- ES is subadditive:

$$\overline{\text{ES}}_p(S_d) = \sum_{i=1}^d \text{ES}_p(X_i).$$

Similarly $\underline{\text{LES}}_p(S_d) = \sum_{i=1}^d \underline{\text{LES}}_p(X_i)$.

- $\overline{\text{VaR}}_p(S_d)$, $\underline{\text{VaR}}_p(S_d)$ and $\underline{\text{ES}}_p(S_d)$: generally open questions

Challenge for $\underline{\text{ES}}_p(S_d)$

To calculate $\underline{\text{ES}}_p(S_d)$ one naturally seeks a **safest** risk in S_d .

Mathematical difficulty

Common understanding of the **most dangerous** scenario:

- Comonotonicity - well accepted notion

Understanding concerning the **safest** scenario:

- $d = 2$: counter-monotonicity
- $d \geq 3$: question mark! (?!)
 - Calls for notions of extremal negative dependence.

Mathematical difficulty

ES respects **convex order**: the natural order of risk preference.

Convex order

We write $X \leq_{cx} Y$ if $\mathbb{E}[f(X)] \leq \mathbb{E}[f(Y)]$ for all convex functions f such that the two expectations exist.

Finding $\underline{ES}_p(S_d)$

Search for a **smallest element** in \mathcal{S}_d with respect to convex order, if it exists.

Mathematical difficulty

VaR does not respect convex order: more tricky

- Good news: the questions for $\overline{\text{VaR}}_p(S_d)$, $\underline{\text{VaR}}_p(S_d)$ and $\underline{\text{ES}}_p(S_d)$ are mathematically similar.

Finding $\overline{\text{VaR}}_p(S_d)$

Search for a **smallest element** in $\mathcal{S}_d(\hat{F}_1, \dots, \hat{F}_d)$ with respect to convex order, where \hat{F}_i is the p -tail-conditional distribution of F_i .

- $\underline{\text{VaR}}_p(S_d)$ is symmetric to $\overline{\text{VaR}}_p(S_d)$.

Summary of existing results

$d = 2$:

- fully solved analytically

$d \geq 3$:

- Homogeneous model ($F_1 = \dots = F_d$)
 - $\underline{ES}_p(S_d)$ solved analytically for decreasing densities, e.g. Pareto, Exponential
 - $\overline{VaR}_p(S_d)$ solved analytically for tail-decreasing densities, e.g. Pareto, Gamma, Log-normal
- Inhomogeneous model
 - Few analytical results: current research
- Numerical methods available: Rearrangement Algorithm

VaR bounds

$d = 2$, Makarov (1981) and Rüschendorf (1982)

For any $p \in (0, 1)$,

$$\overline{\text{VaR}}_p(S_2) = \inf_{x \in [0, 1-p]} \{F_1^{-1}(p+x) + F_2^{-1}(1-x)\},$$

and

$$\underline{\text{VaR}}_p(S_2) = \sup_{x \in [0, p]} \{F_1^{-1}(x) + F_2^{-1}(p-x)\}.$$

- A **large outcome** is coupled with a **small outcome**.

VaR bounds - homogeneous model

Sharp VaR bounds (Wang, Peng and Yang, 2013)

Suppose that the density function of F is decreasing on $[b, \infty)$ for some $b \in \mathbb{R}$. Then, for $p \in [F(b), 1)$, and $X \stackrel{d}{\sim} F$,

$$\overline{\text{VaR}}_p(S_d) = d\mathbb{E}[X|X \in [F^{-1}(p + (d-1)c), F^{-1}(1-c)]],$$

where c is the smallest number in $[0, \frac{1}{d}(1-p)]$ such that

$$\int_{p+(d-1)c}^{1-c} F^{-1}(t) dt \geq \frac{1-p-dc}{d} ((d-1)F^{-1}(p + (d-1)c) + F^{-1}(1-c)).$$

Red part clearly has an ES-type form.

- $c = 0$: $\overline{\text{VaR}}_p(S_d) = \overline{\text{ES}}_p(S_d)$.

VaR bounds - homogeneous model

Sharp VaR bounds II

Suppose that the density function of F is decreasing on its support. Then for $p \in (0, 1)$ and $X \stackrel{d}{\sim} F$,

$$\underline{\text{VaR}}_p(S_d) = \max\{(d-1)F^{-1}(0) + F^{-1}(p), d\mathbb{E}[X|X \leq F^{-1}(p)]\}.$$

Red part has an LES form.

ES bounds - homogeneous model

Sharp ES bounds (Bernard, Jiang and Wang, 2014)

Suppose that the density function of F is decreasing on its support. Then for $p \in (1 - dc, 1)$, $q = (1 - p)/d$ and $X \stackrel{d}{\sim} F$,

$$\begin{aligned}\underline{\text{ES}}_p(S_d) &= \frac{1}{q} \int_0^q \left((d-1)F^{-1}((d-1)t) + F^{-1}(1-t) \right) dt, \\ &= (d-1)^2 \text{LES}_{(d-1)q}(X) + \text{ES}_{1-q}(X),\end{aligned}$$

where c is the smallest number in $[0, \frac{1}{d}]$ such that

$$\int_{(d-1)c}^{1-c} F^{-1}(t) dt \geq \frac{1-dc}{d} \left((d-1)F^{-1}((d-1)c) + F^{-1}(1-c) \right).$$

- One **large outcome** is coupled with $d - 1$ **small outcomes**.

Complete mixability

The homogeneous VaR and ES bounds are based on the notion of **complete mixability**:

Complete mixability, Wang and Wang (2011)

A distribution function F on \mathbb{R} is called d -completely mixable (d -CM) if there exist d random variables $X_1, \dots, X_d \sim F$ such that

$$\mathbb{P}(X_1 + \dots + X_d = dk) = 1,$$

for some $k \in \mathbb{R}$.

- Equivalently, $\mathcal{S}_d(F, \dots, F)$ contains a constant.

Complete mixability

- Some examples of d -CM distributions for all $d \geq 2$:
Normal, Student t, Cauchy, Uniform.
- Most relevant result: F has a **monotone density on a finite interval** with a mean condition (depends on d) is d -CM.
 - Examples: (truncated) Pareto, Gamma, Log-normal.
- Inhomogeneous version called **joint mixability**.
- A full characterization of these classes is at the moment is widely open.

Numerical calculation

Rearrangement Algorithm (RA): Embrechts, Puccetti and Rüschendorf (2013).

- A fast numerical procedure
- Based on the CM-idea
- Discretization of relevant quantile regions
- d possibly large
- Applicable to $\overline{\text{VaR}}_p$, $\underline{\text{VaR}}_p$ and $\underline{\text{ES}}_p$

Asymptotic equivalence

Consider the case $d \rightarrow \infty$. What would happen to $\overline{\text{VaR}}_p(S_d)$?

- Clearly always $\overline{\text{VaR}}_p(S_d) \leq \overline{\text{ES}}_p(S_d)$.
- Recall that $\overline{\text{VaR}}_p(S_d)$ has an ES-type part.

Under some weak conditions,

$$\lim_{d \rightarrow \infty} \frac{\overline{\text{ES}}_p(S_d)}{\overline{\text{VaR}}_p(S_d)} = 1.$$

This was shown first for homogeneous models and then extended to general inhomogeneous models.

Asymptotic equivalence - homogeneous model

Theorem 1

In the homogeneous model, $F_1 = F_2 = \dots = F$, for $p \in (0, 1)$ and $X \sim F$, we have that

$$\lim_{d \rightarrow \infty} \frac{1}{d} \overline{\text{VaR}}_p(S_d) = \text{ES}_p(X).$$

- Similar limits hold for a large class of risk measures

Asymptotic equivalence - worst-cases

Theorem 2 (Embrechts, Wang and Wang, 2014)

Suppose the continuous distributions F_i , $i \in \mathbb{N}$ satisfy that for $X_i \sim F_i$ and some $p \in (0, 1)$,

- (i) $\mathbb{E}[|X_i - \mathbb{E}[X_i]|^k]$ is uniformly bounded for some $k > 1$;
- (ii) $\liminf_{d \rightarrow \infty} \frac{1}{d} \sum_{i=1}^d \text{ES}_p(X_i) > 0$.

Then as $d \rightarrow \infty$,

$$\frac{\overline{\text{ES}}_p(S_d)}{\overline{\text{VaR}}_p(S_d)} = 1 + O(d^{1/k-1}).$$

Asymptotic equivalence - best-cases

Similar results holds for $\underline{\text{VaR}}_p$ and $\underline{\text{ES}}_p$: assume (i) and

$$(iii) \liminf_{d \rightarrow \infty} \frac{1}{d} \sum_{i=1}^d \text{LES}_p(X_i) > 0,$$

then

$$\lim_{d \rightarrow \infty} \frac{\underline{\text{VaR}}_p(S_d)}{\underline{\text{LES}}_p(S_d)} = 1,$$

$$\lim_{d \rightarrow \infty} \frac{\underline{\text{ES}}_p(S_d)}{\sum_{i=1}^d \mathbb{E}[X_i]} = 1,$$

and

$$\frac{\underline{\text{VaR}}_p(S_d)}{\underline{\text{ES}}_p(S_d)} \approx \frac{\sum_{i=1}^d \text{LES}_p(X_i)}{\sum_{i=1}^d \mathbb{E}[X_i]} \leq 1, \quad d \rightarrow \infty.$$

Example: Pareto(2) risks

Bounds on VaR and ES for the sum of d Pareto(2) distributed rvs for $p = 0.999$; VaR_p^+ corresponds to the comonotonic case.

	$d = 8$	$d = 56$
$\underline{\text{VaR}}_p$	31	53
$\underline{\text{ES}}_p$	178	472
VaR_p^+	245	1715
$\overline{\text{VaR}}_p$	465	3454
$\overline{\text{ES}}_p$	498	3486
$\overline{\text{VaR}}_p / \text{VaR}_p^+$	1.898	2.014
$\overline{\text{ES}}_p / \overline{\text{VaR}}_p$	1.071	1.009

Example: Pareto(θ) risks

Bounds on the VaR and ES for the sum of $d = 8$
Pareto(θ)-distributed rvs for $p = 0.999$.

	$\theta = 1.5$	$\theta = 2$	$\theta = 3$	$\theta = 5$	$\theta = 10$
$\overline{\text{VaR}}_p$	1897	465	110	31.65	9.72
$\overline{\text{ES}}_p$	2392	498	112	31.81	9.73
$\overline{\text{ES}}_p / \overline{\text{VaR}}_p$	1.261	1.071	1.018	1.005	1.001

Dependence-uncertainty spread

Theorem 3 (Embrechts, Wang and Wang, 2014)

Take $1 > q \geq p > 0$. Suppose that the continuous distributions F_i , $i \in \mathbb{N}$, satisfy (i) and (iii), and $\limsup_{d \rightarrow \infty} \frac{\sum_{i=1}^d \mathbb{E}[X_i]}{\sum_{i=1}^d \text{ES}_p(X_i)} < 1$, then

$$\liminf_{d \rightarrow \infty} \frac{\overline{\text{VaR}}_q(S_d) - \underline{\text{VaR}}_q(S_d)}{\overline{\text{ES}}_p(S_d) - \underline{\text{ES}}_p(S_d)} \geq 1.$$

- The **uncertainty spread** of VaR is generally bigger than that of ES.
- In recent Consultative Documents of the Basel Committee, $\text{VaR}_{0.99}$ is compared with $\text{ES}_{0.975}$: $p = 0.975$ and $q = 0.99$.

Dependence-uncertainty spread

ES and VaR of $S_d = X_1 + \dots + X_d$, where

- $X_i \sim \text{Pareto}(2 + 0.1i)$, $i = 1, \dots, 5$;
- $X_i \sim \text{Exp}(i - 5)$, $i = 6, \dots, 10$;
- $X_i \sim \text{Log-Normal}(0, (0.1(i - 10))^2)$, $i = 11, \dots, 20$.




	$d = 5$			$d = 20$		
	best	worst	spread	best	worst	spread
$\text{ES}_{0.975}$	22.48	44.88	22.40	29.15	102.35	73.20
$\text{VaR}_{0.975}$	9.79	41.46	31.67	21.44	100.65	79.21
$\text{VaR}_{0.9875}$	12.06	56.21	44.16	22.12	126.63	104.51
$\text{VaR}_{0.99}$	12.96	62.01	49.05	22.29	136.30	114.01
$\frac{\text{ES}_{0.975}}{\text{VaR}_{0.975}}$		1.08			1.02	

Challenges




Open mathematical questions:

- Characterization of complete and joint mixability
- Characterization of \mathcal{S}_d
- Find $\overline{\text{VaR}}_p$ under more general settings, especially in the inhomogeneous model
- Partial dependence information and realistic scenarios
- Marginal uncertainty and statistical estimation
- Many more ...



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THANK YOU!