

Spread-Based Credit Risk Models

Paul Embrechts

London School of Economics
Department of Accounting and Finance

AC 402
FINANCIAL RISK ANALYSIS

Part II

Lent Term, 2003

In this Unit You Will Learn. . .

- the pricing building blocks
 - ★ unconditional payoffs
 - ★ payoffs in survival
 - ★ payoffs at default
- how to price standard instruments: coupon bonds, CDS
- how to extract from the credit spread curve:
 - ★ implied default and survival probabilities
 - ★ conditional survival probabilities
- the calibration of a credit spread curve

What are We Going to Do?

We are looking for a probability distribution for the time of default which is consistent with observed market prices.

There are many such possible probability distributions.

Here, the modeller's / user's intuition is needed, to eliminate implausible default probability distributions.

Therefore we need a setup where it is easy to build an intuition. A setup where the model quantities have a direct relationship to real-world quantities that are known and on which we have experience (and therefore also an opinion).

Mathematically, we are choosing a martingale measure in an incomplete market.

Implied Probabilities

We use:

- Default-free zero coupon bond (ZCB) curve
- Defaultable zero coupon bond (ZCB) curve
- risk-neutral probabilities
- independence of:
 - ★ [defaults and credit spreads dynamics]
 - ★ [interest-rate dynamics]

Default-Free Zero Coupon Bonds:

Payoff: = 1 at time T

Value: = discounted expected payoff

$$B(t, T) = \mathbf{E} [\beta_{t,T} \cdot 1] = \mathbf{E} [\beta_{t,T}] .$$

$\beta_{t,T}$ is the **discount factor** over $[t, T]$.

$$\beta_{t,T} = \exp\left\{-\int_t^T r(s)ds\right\}$$

where $r(s)$ is the continuously compounded short rate at time t .

Defaultable Zerobonds:

$$\text{Payoff} = \mathbf{1}_{\{\tau > T\}} = \begin{cases} 1 & \text{if default after } T, \text{ i.e. } \tau > T, \\ 0 & \text{if default before } T, \text{ i.e. } \tau \leq T. \end{cases}$$

Value = discounted expected payoff

$$\overline{B}(t, T) = \mathbf{E} \left[\beta_{t, T} \mathbf{1}_{\{\tau > T\}} \right].$$

$\mathbf{1}_{\{A\}}$ is the indicator function of event A :

$\mathbf{1}_{\{A\}} = 1$ if A is true, and $\mathbf{1}_{\{A\}} = 0$ if A is false.

-

If there is no correlation:

$$\begin{aligned}\overline{B}(t, T) &= \mathbf{E} \left[\beta_{t, T} \mathbf{1}_{\{\tau > T\}} \right] \\ &= \mathbf{E} \left[\beta_{t, T} \right] \mathbf{E} \left[\mathbf{1}_{\{\tau > T\}} \right] \\ &= B(t, T) \mathbf{E} \left[\mathbf{1}_{\{\tau > T\}} \right] \\ &= B(t, T) P(t, T)\end{aligned}$$

$P(t, T)$ is the implied probability of survival in $[t, T]$.

($\beta_{t, T}$ is the discount-factor over $[t, T]$.)

The Implied Survival Probability

the implied survival probability is the ratio of the ZCB prices:

$$P(t, T) = \frac{\overline{B}(t, T)}{B(t, T)}$$

- Default Probability = 1- Survival Probabilities
- initially at one

$$P(t, t) = 1$$

- eventually there is a default:

$$P(t, \infty) = 0$$

- $P(t, T)$ is decreasing in T .

Conditional Survival Probabilities

Want to focus on *one time interval in the future*.

Probability of survival in $[T_1, T_2]$. . .

. . . given that there was no default until T_1 .

$$P(t, T_1, T_2) = \frac{P(t, T_2)}{P(t, T_1)} = \frac{\overline{B}(t, T_2) B(t, T_1)}{B(t, T_2) \overline{B}(t, T_1)}.$$

- Survival to T = survival to s AND survival from s to T

$$P(t, T) = P(t, s)P(t, s, T)$$

- $P(t, T, T) = 1$, unless default is 'scheduled' at T .

Note: Bayes' Rule:

$$\Pr[A|B] = \frac{\Pr[A \cap B]}{\Pr[B]}.$$

Here: A = survival until T_2 B = survival until T_1

$A \cap B$ = [survival until T_2] AND [survival until T_1]
= [survival until T_2]

Numerical Example

Mat	Yield (%)		ZCB		Probabilities (%)		
	Gov	Issuer	Gov	Issuer	Surv.	CSurv.	CDef/T
0,5	5,75	7	0,9724	0,9667	99,41	99,41	1,17
1	6,10	7,85	0,9425	0,9272	98,38	98,96	2,09
3	6,25	8,25	0,8337	0,7883	94,56	96,12	1,94
5	6,40	8,65	0,7333	0,6605	90,07	95,25	2,38
7	6,78	9,08	0,6318	0,5442	86,14	95,64	2,18
10	6,95	9,70	0,5107	0,3962	77,58	90,06	3,31

Numbers taken from Das (1997)

Forward Rates

define the *default-free forward rate* $F_k(t)$, the *defaultable forward rate* $\bar{F}_k(t)$ and the *discrete forward default intensity* $H_k(t)$ over $[T_k, T_{k+1}]$ as seen from t :

$$F_k(t) = \frac{1}{\delta_k} \left(\frac{B_k(t)}{B_{k+1}(t)} - 1 \right).$$

$$\bar{F}_k(t) = \frac{1}{\delta_k} \left(\frac{\bar{B}_k(t)}{\bar{B}_{k+1}(t)} - 1 \right).$$

$$H_k(t) = \frac{1}{\delta_k} \left(\frac{P_k(t)}{P_{k+1}(t)} - 1 \right).$$

where $\delta_k = T_{k+1} - T_k$ and $P_k(t) = P(t, T_k)$.

Relation to Forward Spreads

the **conditional probability of default** over $[T_k, T_{k+1}]$ is given by:

$$\frac{1}{\delta_k} P^{\text{def}}(t, T_k, T_{k+1}) = \frac{\bar{F}_k(t) - F_k(t)}{1 + \delta_k \bar{F}_k(t)} = \frac{H_k(t)}{1 + \delta_k H_k(t)}$$

[Default Probability]

= [Length of time interval] \times [Spread of forward rates]

\times [Discounting with defaultable forward rate]

Local Default Probabilities

use **continuously compounded forward rates**:

$$f(t, T) = \lim_{\Delta t \searrow 0} F(t, T, T + \Delta t) = -\frac{\partial}{\partial T} \ln B(t, T)$$

$$\bar{f}(t, T) = \lim_{\Delta t \searrow 0} \bar{F}(t, T, T + \Delta t) = -\frac{\partial}{\partial T} \ln \bar{B}(t, T)$$

The **local probability of default** at time T is:

$$\lim_{\Delta t \searrow 0} \frac{P^{\text{def}}(t, T, T + \Delta t)}{\Delta t} = \bar{f}(t, T) - f(t, T)$$

The probability of default in $[T, T + \Delta t]$ is approximately **proportional** to the length of the interval $[T, T + \Delta t]$ with proportionality factor $(\bar{f}(t, T) - f(t, T))$

Why Poisson Processes?

Ultimate goal:

A mathematical model of defaults that is realistic and tractable and useful for pricing and hedging.

Defaults are

- sudden, usually unexpected
- rare (hopefully :-)
- cause large, *discontinuous* price changes.

Require from the mathematical model the same properties.

Furthermore: Previous section

The probability of default in a short time interval is approximately **proportional** to the length of the interval.

What is a Poisson Process?

$N(t)$ = value of the process at time t .

- Starts at zero: $N(0) = 0$
- Integer-valued: $N(t) = 0, 1, 2, \dots$
- Increasing or constant
- Main use: marking points in time
 T_1, T_2, \dots the jump times of N
- Here **Default**: time of the first jump of N
 $\tau = T_1$
- Jump probability over small intervals proportional to that interval.

- Proportionality factor = λ

BTW: Except for the last two points, the same notation and properties apply to *Point Processes*, too.

Discrete-time approximation:

- divide $[0, T]$ in n intervals of equal length

$$\Delta t = T/n$$

- Make the jump probability in each interval $[t_i, t_i + \Delta t]$ proportional to Δt :

$$p := \mathbf{P} [N(t_i + \Delta t) - N(t_i) = 1] = \lambda \Delta t.$$

(these are independent across intervals.)

- more exact approximation: $p = 1 - e^{-\lambda \Delta t}$
- Let $n \rightarrow \infty$ or $\Delta t \rightarrow 0$.

Important Properties

Homogeneous Poisson process with intensity λ

Jump Probabilities over interval $[t, T]$:

- No jump:

$$\mathbf{P} [N(T) = N(t)] = \exp\{-(T - t)\lambda\}$$

- n jumps:

$$\mathbf{P} [N(T) = N(t) + n] = \frac{1}{n!} (T - t)^n \lambda^n e^{-(T-t)\lambda}.$$

- Inter-arrival times $\mathbf{P} [(T_{n+1} - T_n) \in t dt] = \lambda e^{-\lambda t} dt.$
- Expectation (locally) $\mathbf{E} [dN] = \lambda dt.$

Distribution of the Time of the first Jump

T_1 time of first jump.

Distribution: $F(t) := \mathbf{P} [T_1 \leq t]$

Know probability of no jump until T :

$$\mathbf{P} [N(T) = 0] = e^{-\lambda T}$$

= probability of $T_1 > T$. Thus

$$1 - F(t) = e^{-\lambda t}$$

$$F(t) = 1 - e^{-\lambda t}$$

$$F'(t) = f(t) = \lambda e^{-\lambda t}.$$

-

- ⇒ T_1 is exponentially distributed with parameter λ .
- ⇒ This is also the distribution of the *next* jump, given that there have been k jumps so far.
- ⇒ Independently of how much time has passed so far:
It never is 'about time a jump happened'
or 'nothing has happened, I don't think anything will happen any more. . . '

Inhomogeneous Poisson Process

Inhomogeneous = with **time-dependent** intensity function $\lambda(t)$

Probability of no jumps (survival):

$$\mathbf{P} [N_T = N_t] = \exp \left\{ - \int_t^T \lambda(s) ds \right\}.$$

Probability of n jumps:

$$\mathbf{P} [N_T - N_t = n] = \frac{1}{n!} \left(\int_t^T \lambda(s) ds \right)^n e^{- \int_t^T \lambda(s) ds}.$$

Density of the time of the first jump:

$$\begin{aligned}\mathbf{P} [T_1 \in [a, b]] &= \int_a^b f(t, u) du \\ &= \int_a^b \lambda(u) e^{-\int_t^u \lambda(s) ds} du.\end{aligned}$$

Cox Processes

default rate = Intensity of PP = credit spread.

Credit spreads are stochastic.  Need stochastic intensity

- define a stochastic intensity process λ , e.g.

$$d\lambda = \mu_\lambda dt + \sigma_\lambda dW$$

- $\lambda(t)\Delta t$: default probability over the next time-interval $[t, t + \Delta t]$.
(That's all we need to know at t .)
- at $t + \Delta t$: Intensity has changed, $\lambda(t + \Delta t) = \lambda(t) + d\lambda$ is new (local) default probability.
- Conditional on the realisation of the intensity process, the Cox process is an inhomogeneous Poisson process.

Cox Processes: The Conditioning-Trick

The Gods are gambling in a certain sequence:

- First, the full path of the intensity $\lambda(t)$ is drawn from all possible paths for $\lambda(t)$.
- Then they take this $\lambda(t)$ and *use it as intensity for an inhomogeneous Poisson process N* .

They draw the jumps of $N(t)$ according to this distribution.

- Then the information is revealed to the mortals:
At time t they may only know $\lambda(s)$ and $N(s)$ for s up to t .

The Conditioning Trick

First, pretend you knew the path of λ , what would the price be? (Depending on λ , of course.)

Then average over the possible paths of λ .

Let $X(N)$ be a payoff that we want to value. It depends on the question whether a default occurred ($N = 1$) or not ($N = 0$).

$$\mathbf{E} [X(N)] = \mathbf{E} [\mathbf{E} [X(N) \mid \lambda(t) \ \forall t]]$$

The inner expectation is easily calculated treating N as an inhomogeneous Poisson Process.

Properties of Cox Processes

Probability of no jumps (survival):

$$\mathbf{P} [N_T = N_t] = \mathbf{E} \left[\exp \left\{ - \int_t^T \lambda(s) ds \right\} \right].$$

Probability of n jumps:

$$\mathbf{P} [N_T - N_t = n] = \mathbf{E} \left[\frac{1}{n!} \left(\int_t^T \lambda(s) ds \right)^n e^{- \int_t^T \lambda(s) ds} \right].$$

Density of the time of the first jump:

$$\mathbf{P} [T_1 \in [a, b]] = \int_a^b f(t, u) du$$

$$= \int_a^b \mathbf{E} \left[\lambda(u) e^{-\int_t^u \lambda(s) ds} \right] du.$$

The expectations always only refer to the realisation of λ .

Models using Poisson-type Default Times

Jarrow / Turnbull (JoF 50(1), p. 53ff.):
default triggered by constant intensity Poisson Process

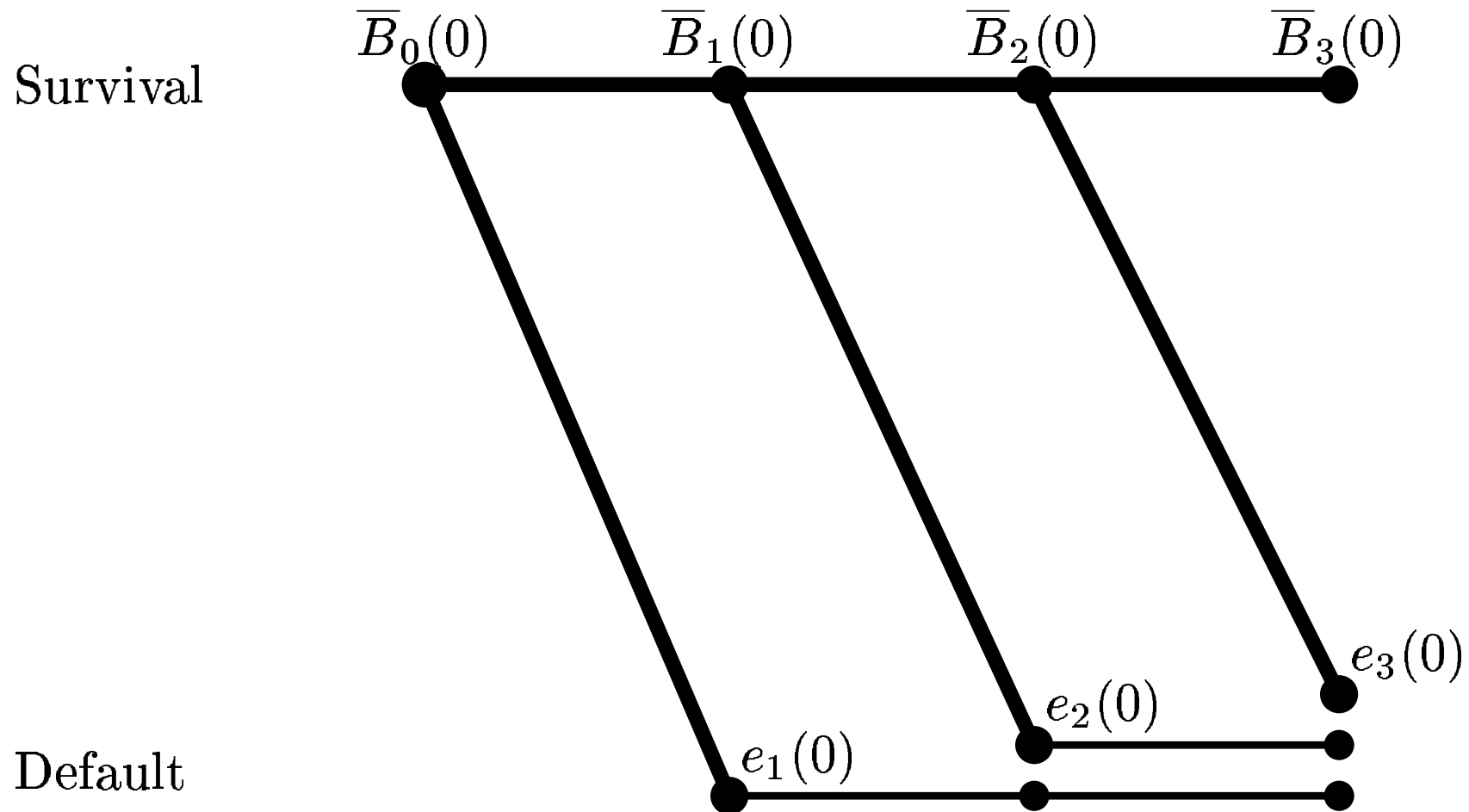
Duffie / Singleton (JoF 52, pp. 1287 ff., 1997):
stochastic-intensity default trigger

Madan / Unal: (Review of Derivatives Research 2, pp. 121 ff., 1998)
Default intensity depends on firm's value

Jarrow / Lando / Turnbull:
default intensity is a function of the rating. The rating follows a Markov chain.

Schönbucher (1998, 1999):
stochastic intensity, modelling a full term-structure of future default intensities

A Simple Tree Model



Setup of the Basic Model

discrete points in time

$$0 = T_0, T_1, T_2, \dots, T_N$$

converge to continuous-time when distances $\rightarrow 0$

- $B(t, T_k)$ or $B_k(t)$: default-free zerobond price with maturity T_k
- $\bar{B}(t, T_k)$ or $\bar{B}_k(t)$: defaultable zerobond price with maturity T_k
- τ : time of default

The conditional survival probability is the “across” branching probability here.

Price Building Blocks

Payoffs *independent of default* (possibly stochastic)

- 'classical' pricing problem
- building block

$B_k(t)$ default-free zero coupon bond with maturity T_k

Payoffs *in survival* (possibly stochastic)

- zero recovery (we get the payoff *only* in survival)
- building block

$\bar{B}_k(t)$ defaultable ZCB, maturity T_k , zero recovery

- \bar{B} -prices follow directly from defaultable forward rates \bar{F}_k (defined later)

-

Payoffs *at default*

- for example recoveries
- building block:
 $e_k(t)$ price at time t of \$1 at default iff default in $[T_{k-1}, T_k]$
- prices $e_k(t)$ are already fully determined by the prices and dynamics of B_k and \overline{B}_k

Pricing the Building Blocks $e_k(t)$

Payoff of $e_k = 1$ at T_k if and only if default in $[T_{k-1}, T_k]$

- Probability of default in $[T_{k-1}, T_k]$

$$= P_{k-1}(t) - P_k(t) = \delta_{k-1} H_{k-1}(t) P_k(t)$$

- Discounting until T_k using $B_k(t)$ yields

$$e_k(t) = \delta_{k-1} H_{k-1}(t) \bar{B}_k(t)$$

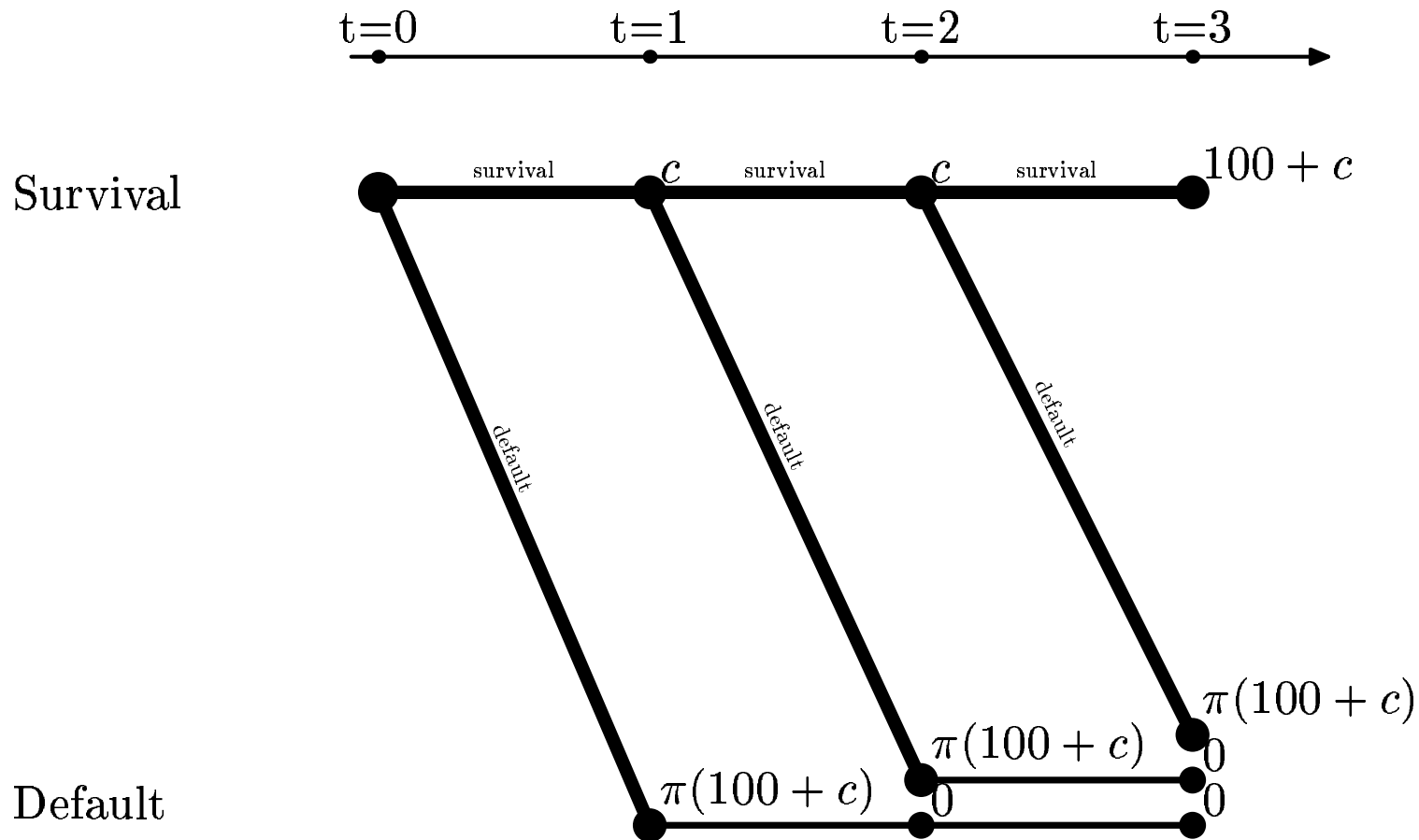
(remember $P_k B_k = \bar{B}_k$)

Here we used independence again. Pricing is more complicated otherwise. (See approximative solutions in paper.)

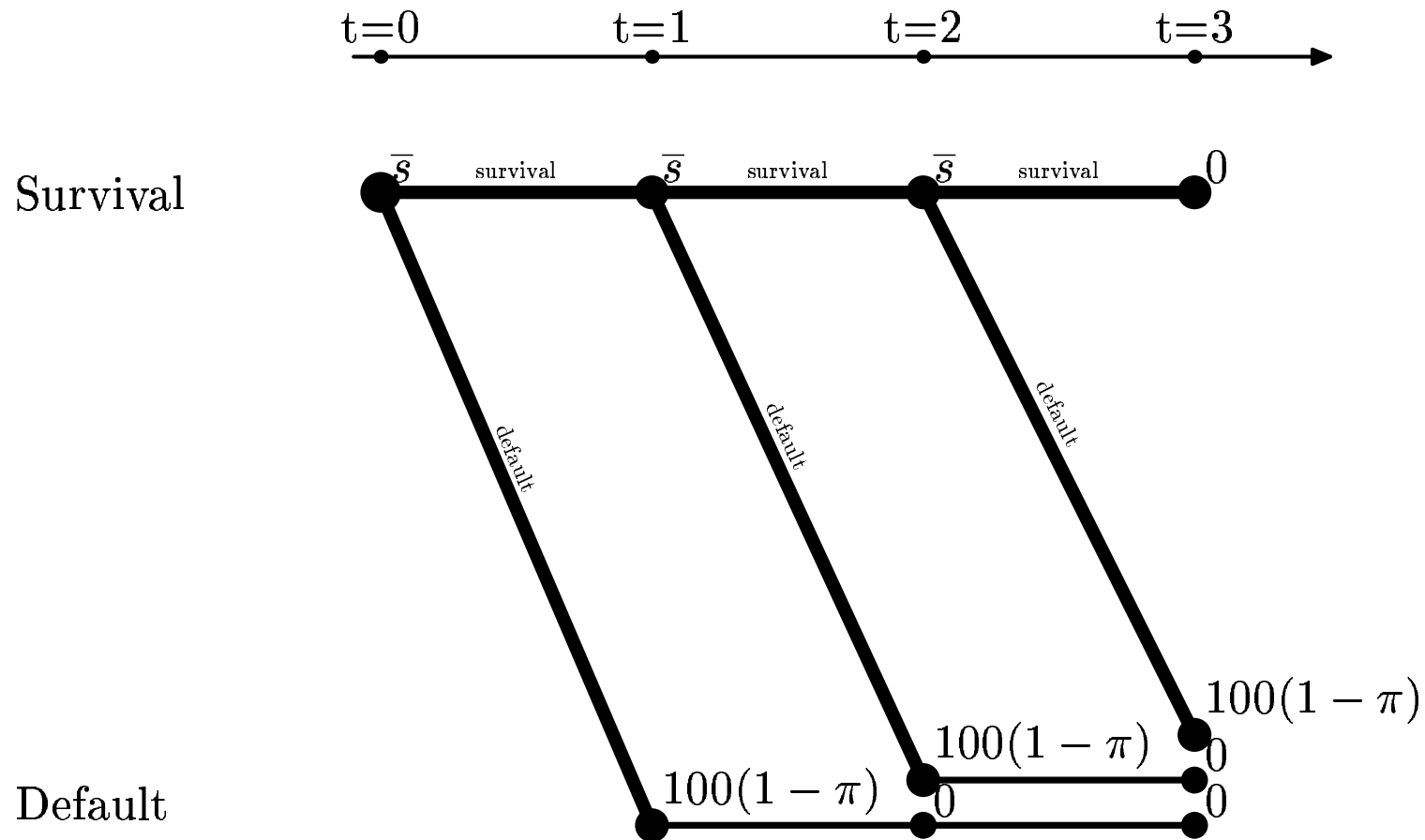
Pricing Defaultable Securities

- Identify the payoffs in events of survival and default
- Payoffs in default:
 - ★ recovery rate π
 - ★ take care specifying the claim size in default (notional, accrued interest, amortisation)
 - ★ observe seniority
- For calibration securities this should be straightforward.
- Model price = weighted sum of building-block prices

Pricing a Defaultable Coupon Bond



Pricing a Credit Default Swap



Pricing with Building Blocks

Fixed Coupon Bonds: coupon c , payable at T_i , recovery π

$$\underbrace{\sum_{i=1}^N c \bar{B}_i}_{\text{coupons}} \quad \underbrace{+ \bar{B}_N}_{+\text{principal}} \quad \underbrace{+ \pi(1+c) \sum_{i=1}^N e_i}_{+\text{recovery}}$$

Credit Default Swap: fee s , payable at T_i , payoff at default π

$$\underbrace{- \sum_{i=1}^N s \bar{B}_i}_{-\text{fee}} \quad \underbrace{+ \pi \sum_{i=1}^N e_i}_{+\text{default payoff}}$$

Exercise: Amortisation, Floating-Rate Notes

What have we reached?

- Prices of defaultable securities are given in terms of the building blocks B_k , \bar{B}_k and e_k
- The prices of the building blocks can be represented in terms of “forward-rates” F_k and H_k
- Fitting default-free forward rates F_k is standard.
- Can adapt forward rate - fitting methods to fit H_k

References

- [1] Philippe Artzner and Freddy Delbaen. Credit risk and prepayment option. *Astin Bulletin*, 22:81–96, 1992.
- [2] Philippe Artzner and Freddy Delbaen. Default risk insurance and incomplete markets. Working paper, Université Luis Pasteur et C.N.R.S., Strasbourg, June 1994.
- [3] Pierre Brémaud. *Point Processes and Queues*. Springer, Berlin, Heidelberg, New York, 1981.
- [4] Sanjiv Ranjan Das. Credit risk derivatives. *Journal of Derivatives*, 2:7–23, 1995.
- [5] Gregory R. Duffee. Estimating the price of default risk. Working paper, Federal Reserve Board, Washington DC, September 1995.
- [6] Darrel Duffie and Kenneth Singleton. An econometric model of the term structure of interest rate swap yields. *Journal of Finance*, 52(4):1287–1321, 1997.
- [7] Darrell Duffie. Forward rate curves with default risk. Working paper, Graduate School of Business, Stanford University, December 1994.
- [8] Darrell Duffie and Ming Huang. Swap rates and credit quality. Working paper, Graduate School of Business, Stanford University, October 1994.
- [9] Darrell Duffie, Mark Schroder, and Costis Skiadas. Recursive valuation of defaultable securities and the timing of resolution of uncertainty. Working Paper 195, Kellogg Graduate School of Management, Department of Finance, Northwestern University, November 1994. revised March 1995.
- [10] Darrell Duffie and Kenneth J. Singleton. Econometric modeling of term structures of defaultable bonds. Working paper, Graduate School of Business, Stanford University, June 1994. revised November 1994.
- [11] Bjorn Flesaker, Lane Houghston, Laurence Schreiber, and Lloyd Sprung. Taking all the credit. *Risk Magazine*, 7:105–108, 1994.
- [12] Robert A. Jarrow, David Lando, and Stuart M. Turnbull. A Markov model for the term structure of credit risk spreads. Working paper, Johnson Graduate School of Management, Cornell University, December 1993. revised January 1994.
- [13] Robert A. Jarrow and Stuart M. Turnbull. Pricing derivatives on financial securities subject to credit risk. *Journal of Finance*, 50:53–85, 1995.

- [14] David Lando. On Cox processes and credit risky bonds. Working paper, Institute of Mathematical Statistics, University of Copenhagen, March 1994. revised December 1994.
- [15] Dilip B. Madan and Haluk Unal. Pricing the risks of default. Working paper, College of Business and Management, University of Maryland, March 1994.
- [16] Philipp J. Schönbucher. The term structure of defaultable bond prices. Discussion Paper B-384, University of Bonn, SFB 303, August 1996.
- [17] Philipp J. Schönbucher. Modelling defaultable bond prices. Working paper, London School of Economics, Financial Markets Group, July 1997.
- [18] Philipp J. Schönbucher. Pricing credit risk derivatives. Working Paper wp-10, London School of Economics, Financial Markets Group, July 1997.